# Lecture 2: From Linear Regression to Kalman Filter and Beyond

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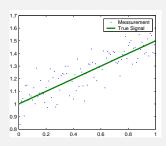
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# Batch Linear Regression [1/2]



Consider the linear regression model

$$y_k = a_1 + a_2 t_k + \epsilon_k,$$

with 
$$\epsilon_k \sim N(0, \sigma^2)$$
 and  $\mathbf{a} = (a_1, a_2) \sim N(\mathbf{m}_0, \mathbf{P}_0)$ .

In probabilistic notation this is:

$$p(y_k \mid \mathbf{a}) = N(y_k \mid \mathbf{H}_k \mathbf{a}, \sigma^2)$$
$$p(\mathbf{a}) = N(\mathbf{a} \mid \mathbf{m}_0, \mathbf{P}_0),$$

where **H**<sub>*k*</sub> = (1  $t_k$ ).

# Batch Linear Regression [2/2]

The Bayesian batch solution by the Bayes' rule:

$$p(\mathbf{a} \mid y_{1:N}) \propto p(\mathbf{a}) \prod_{k=1}^{N} p(y_k \mid \mathbf{a})$$

$$= N(\mathbf{a} \mid \mathbf{m}_0, \mathbf{P}_0) \prod_{k=1}^{N} N(y_k \mid \mathbf{H}_k \mathbf{a}, \sigma^2).$$

The posterior is Gaussian

$$p(\mathbf{a} \mid y_{1:N}) = N(\mathbf{a} \mid \mathbf{m}_N, \mathbf{P}_N).$$

• The mean and covariance are given as

$$\mathbf{m}_N = \left[ \mathbf{P}_0^{-1} + \frac{1}{\sigma^2} \mathbf{H}^T \mathbf{H} \right]^{-1} \left[ \frac{1}{\sigma^2} \mathbf{H}^T \mathbf{y} + \mathbf{P}_0^{-1} \mathbf{m}_0 \right]$$

$$\mathbf{P}_N = \left[ \mathbf{P}_0^{-1} + \frac{1}{\sigma^2} \mathbf{H}^T \mathbf{H} \right]^{-1},$$
where  $\mathbf{H}_k = (1 \ t_k)$  and  $\mathbf{H} = (\mathbf{H}_1; \mathbf{H}_2; \dots; \mathbf{H}_N)$ , and

# Recursive Linear Regression [1/3]

 Assume that we have already computed the posterior distribution, which is conditioned on the measurements up to k - 1:

$$p(\mathbf{a} \mid y_{1:k-1}) = N(\mathbf{a} \mid \mathbf{m}_{k-1}, \mathbf{P}_{k-1}).$$

• Assume that we get the kth measurement  $y_k$ . Using the equations from the previous slide we get

$$p(\mathbf{a} \mid y_{1:k}) \propto p(y_k \mid \mathbf{a}) p(\mathbf{a} \mid y_{1:k-1})$$
$$\propto N(\mathbf{a} \mid \mathbf{m}_k, \mathbf{P}_k).$$

• The mean and covariance are given as

$$\mathbf{m}_{k} = \left[\mathbf{P}_{k-1}^{-1} + \frac{1}{\sigma^{2}}\mathbf{H}_{k}^{T}\mathbf{H}_{k}\right]^{-1} \left[\frac{1}{\sigma^{2}}\mathbf{H}_{k}^{T}y_{k} + \mathbf{P}_{k-1}^{-1}\mathbf{m}_{k-1}\right]$$
$$\mathbf{P}_{k} = \left[\mathbf{P}_{k-1}^{-1} + \frac{1}{\sigma^{2}}\mathbf{H}_{k}^{T}\mathbf{H}_{k}\right]^{-1}.$$

## Recursive Linear Regression [2/3]

By the matrix inversion lemma (or Woodbury identity):

$$\mathbf{P}_k = \mathbf{P}_{k-1} - \mathbf{P}_{k-1} \mathbf{H}_k^T \left[ \mathbf{H}_k \mathbf{P}_{k-1} \mathbf{H}_k^T + \sigma^2 \right]^{-1} \mathbf{H}_k \mathbf{P}_{k-1}.$$

Now the equations for the mean and covariance reduce to

$$S_k = \mathbf{H}_k \mathbf{P}_{k-1} \mathbf{H}_k^T + \sigma^2$$

$$\mathbf{K}_k = \mathbf{P}_{k-1} \mathbf{H}_k^T S_k^{-1}$$

$$\mathbf{m}_k = \mathbf{m}_{k-1} + \mathbf{K}_k [y_k - \mathbf{H}_k \mathbf{m}_{k-1}]$$

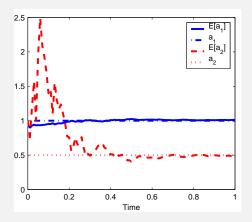
$$\mathbf{P}_k = \mathbf{P}_{k-1} - \mathbf{K}_k S_k \mathbf{K}_k^T.$$

- Computing these for k = 0, ..., N gives exactly the linear regression solution but without a matrix inversion<sup>1</sup>!
- A special case of Kalman filter.

<sup>&</sup>lt;sup>1</sup>Without an *explicit* matrix inversion

## Recursive Linear Regression [3/3]

Convergence of the recursive solution to the batch solution – on the last step the solutions are exactly equal:



## Batch vs. Recursive Estimation [1/2]

#### General batch solution:

Specify the measurement model:

$$p(\mathbf{y}_{1:N} | \boldsymbol{\theta}) = \prod_{k} p(\mathbf{y}_{k} | \boldsymbol{\theta}).$$

- Specify the prior distribution  $p(\theta)$ .
- Compute posterior distribution by the Bayes' rule:

$$p(\theta \mid \mathbf{y}_{1:N}) = \frac{1}{Z}p(\theta) \prod_{k} p(\mathbf{y}_{k} \mid \theta).$$

 Compute point estimates, moments, predictive quantities etc. from the posterior distribution.

# Batch vs. Recursive Estimation [2/2]

#### General recursive solution:

- Specify the measurement likelihood  $p(\mathbf{y}_k | \theta)$ .
- Specify the prior distribution  $p(\theta)$ .
- Process measurements y<sub>1</sub>,..., y<sub>N</sub> one at a time, starting from the prior:

$$p(\theta \mid \mathbf{y}_1) = \frac{1}{Z_1} p(\mathbf{y}_1 \mid \theta) p(\theta)$$

$$p(\theta \mid \mathbf{y}_{1:2}) = \frac{1}{Z_2} p(\mathbf{y}_2 \mid \theta) p(\theta \mid \mathbf{y}_1)$$

$$\vdots$$

$$p(\theta \mid \mathbf{y}_{1:N}) = \frac{1}{Z_N} p(\mathbf{y}_N \mid \theta) p(\theta \mid \mathbf{y}_{1:N-1}).$$

 The posterior at the last step is the same as the batch solution.

# Advantages of Recursive Solution

- The recursive solution can be considered as the online learning solution to the Bayesian learning problem.
- Batch Bayesian inference is a special case of recursive Bayesian inference.
- The parameter can be modeled to change between the measurement steps ⇒ basis of filtering theory.

# Drift Model for Linear Regression [1/3]

 Let assume Gaussian random walk between the measurements in the linear regression model:

$$p(y_k | \mathbf{a}_k) = N(y_k | \mathbf{H}_k \mathbf{a}_k, \sigma^2)$$
  
 $p(\mathbf{a}_k | \mathbf{a}_{k-1}) = N(\mathbf{a}_k | \mathbf{a}_{k-1}, \mathbf{Q})$   
 $p(\mathbf{a}_0) = N(\mathbf{a}_0 | \mathbf{m}_0, \mathbf{P}_0).$ 

Again, assume that we already know

$$p(\mathbf{a}_{k-1} | y_{1:k-1}) = N(\mathbf{a}_{k-1} | \mathbf{m}_{k-1}, \mathbf{P}_{k-1}).$$

• The joint distribution of  $\mathbf{a}_k$  and  $\mathbf{a}_{k-1}$  is (due to Markovianity of dynamics!):

$$p(\mathbf{a}_k, \mathbf{a}_{k-1} | y_{1:k-1}) = p(\mathbf{a}_k | \mathbf{a}_{k-1}) p(\mathbf{a}_{k-1} | y_{1:k-1}).$$

# Drift Model for Linear Regression [2/3]

• Integrating over  $\mathbf{a}_{k-1}$  gives:

$$p(\mathbf{a}_k | y_{1:k-1}) = \int p(\mathbf{a}_k | \mathbf{a}_{k-1}) p(\mathbf{a}_{k-1} | y_{1:k-1}) d\mathbf{a}_{k-1}.$$

- This equation for Markov processes is called the Chapman-Kolmogorov equation.
- Because the distributions are Gaussian, the result is Gaussian

$$p(\mathbf{a}_k | y_{1:k-1}) = N(\mathbf{a}_k | \mathbf{m}_k^-, \mathbf{P}_k^-),$$

where

$$egin{aligned} \mathbf{m}_k^- &= \mathbf{m}_{k-1} \ \mathbf{P}_k^- &= \mathbf{P}_{k-1} + \mathbf{Q}. \end{aligned}$$

# Drift Model for Linear Regression [3/3]

As in the pure recursive estimation, we get

$$p(\mathbf{a}_k \mid y_{1:k}) \propto p(y_k \mid \mathbf{a}_k) p(\mathbf{a}_k \mid y_{1:k-1})$$
$$\propto N(\mathbf{a}_k \mid \mathbf{m}_k, \mathbf{P}_k).$$

 After applying the matrix inversion lemma, mean and covariance can be written as

$$\begin{split} \mathcal{S}_k &= \mathbf{H}_k \mathbf{P}_k^- \mathbf{H}_k^T + \sigma^2 \\ \mathbf{K}_k &= \mathbf{P}_k^- \mathbf{H}_k^T \mathcal{S}_k^{-1} \\ \mathbf{m}_k &= \mathbf{m}_k^- + \mathbf{K}_k [y_k - \mathbf{H}_k \mathbf{m}_k^-] \\ \mathbf{P}_k &= \mathbf{P}_k^- - \mathbf{K}_k \mathcal{S}_k \mathbf{K}_k^T. \end{split}$$

- Again, we have derived a special case of the Kalman filter.
- The batch version of this solution would be much more complicated.

## State Space Notation

In the previous section we formulated the model as

$$p(\mathbf{a}_k \mid \mathbf{a}_{k-1}) = N(\mathbf{a}_k \mid \mathbf{a}_{k-1}, \mathbf{Q})$$
$$p(y_k \mid \mathbf{a}_k) = N(y_k \mid \mathbf{H}_k \mathbf{a}_k, \sigma^2)$$

- But in Kalman filtering and control theory the vector of parameters  $\mathbf{a}_k$  is usually called "state" and denoted as  $\mathbf{x}_k$ .
- More standard state space notation:

$$p(\mathbf{x}_k | \mathbf{x}_{k-1}) = N(\mathbf{x}_k | \mathbf{x}_{k-1}, \mathbf{Q})$$
$$p(y_k | \mathbf{x}_k) = N(y_k | \mathbf{H}_k \mathbf{x}_k, \sigma^2)$$

Or equivalently

$$\mathbf{x}_k = \mathbf{x}_{k-1} + \mathbf{q}$$
  
 $y_k = \mathbf{H}_k \, \mathbf{x}_k + r,$ 

where  $\mathbf{q} \sim N(\mathbf{0}, \mathbf{Q}), r \sim N(0, \sigma^2).$ 

#### Kalman Filter [1/2]

The canonical Kalman filtering model is

$$p(\mathbf{x}_k \mid \mathbf{x}_{k-1}) = N(\mathbf{x}_k \mid \mathbf{A}_{k-1} \mathbf{x}_{k-1}, \mathbf{Q}_{k-1})$$
$$p(\mathbf{y}_k \mid \mathbf{x}_k) = N(\mathbf{y}_k \mid \mathbf{H}_k \mathbf{x}_k, \mathbf{R}_k).$$

More often, this model can be seen in the form

$$\mathbf{x}_{k} = \mathbf{A}_{k-1} \, \mathbf{x}_{k-1} + \mathbf{q}_{k-1}$$
  
 $\mathbf{y}_{k} = \mathbf{H}_{k} \, \mathbf{x}_{k} + \mathbf{r}_{k}.$ 

The Kalman filter actually calculates the following distributions:

$$p(\mathbf{x}_k \mid \mathbf{y}_{1:k-1}) = N(\mathbf{x}_k \mid \mathbf{m}_k^-, \mathbf{P}_k^-)$$
$$p(\mathbf{x}_k \mid \mathbf{y}_{1:k}) = N(\mathbf{x}_k \mid \mathbf{m}_k, \mathbf{P}_k).$$

## Kalman Filter [2/2]

Prediction step of the Kalman filter:

$$\begin{split} & \boldsymbol{m}_k^- = \boldsymbol{A}_{k-1} \, \boldsymbol{m}_{k-1} \\ & \boldsymbol{P}_k^- = \boldsymbol{A}_{k-1} \, \boldsymbol{P}_{k-1} \, \boldsymbol{A}_{k-1}^T + \boldsymbol{Q}_{k-1}. \end{split}$$

Update step of the Kalman filter:

$$\begin{split} \mathbf{S}_k &= \mathbf{H}_k \, \mathbf{P}_k^- \, \mathbf{H}_k^T + \mathbf{R}_k \\ \mathbf{K}_k &= \mathbf{P}_k^- \, \mathbf{H}_k^T \, \mathbf{S}_k^{-1} \\ \mathbf{m}_k &= \mathbf{m}_k^- + \mathbf{K}_k \, [\mathbf{y}_k - \mathbf{H}_k \, \mathbf{m}_k^-] \\ \mathbf{P}_k &= \mathbf{P}_k^- - \mathbf{K}_k \, \mathbf{S}_k \, \mathbf{K}_k^T. \end{split}$$

 These equations will be derived from the general Bayesian filtering equations in the next lecture.

# Probabilistic Non-Linear Filtering [1/2]

Generic discrete-time state space models

$$\mathbf{x}_k = \mathbf{f}(\mathbf{x}_{k-1}, \mathbf{q}_{k-1})$$
  
 $\mathbf{y}_k = \mathbf{h}(\mathbf{x}_k, \mathbf{r}_k).$ 

Generic Markov models

$$\mathbf{y}_k \sim p(\mathbf{y}_k \,|\, \mathbf{x}_k)$$
  
 $\mathbf{x}_k \sim p(\mathbf{x}_k \,|\, \mathbf{x}_{k-1}).$ 

 Approximation methods: Extended Kalman filters (EKF), Unscented Kalman filters (UKF), sequential Monte Carlo (SMC) filters a'ka particle filters.

# Probabilistic Non-Linear Filtering [2/2]

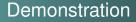
- In continuous-discrete filtering models, dynamics are modeled in continuous time, measurements at discrete time steps.
- The continuous time versions of Markov models are called as stochastic differential equations:

$$\frac{d\mathbf{x}}{dt} = \mathbf{f}(\mathbf{x}, t) + \mathbf{w}(t)$$

- where  $\mathbf{w}(t)$  is a continuous time Gaussian white noise process.
- Approximation methods: Extended Kalman filters, Unscented Kalman filters, sequential Monte Carlo, particle filters.

# Summary

- Linear regression problem can be solved as batch problem or recursively – the latter solution is a special case of Kalman filter.
- A generic Bayesian estimation problem can also be solved as batch problem or recursively.
- If we let the linear regression parameter change between the measurements, we get a simple linear state space model – again solvable with Kalman filtering model.
- By generalizing this idea and the solution we get the Kalman filter algorithm.
- By further generalizing to non-Gaussian models results in a generic probabilistic state space model.



Batch and recursive linear regression.