Lecture 6: Particle Filtering

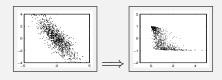
Simo Särkkä

March 9, 2016

Contents

- Principle of Particle Filter
- Monte Carlo Integration and Importance Sampling
- Sequential Importance Sampling and Resampling
- Rao-Blackwellized Particle Filter
- Particle Filter Properties
- 6 Summary and Demonstration

Particle Filtering: Principle



- Animation: Kalman vs. Particle Filtering:
 - Naiman filter animation
 - Particle filter animation
- The idea is to form a weighted particle presentation $(\mathbf{x}^{(i)}, \mathbf{w}^{(i)})$ of the posterior distribution:

$$p(\mathbf{x}) \approx \sum_{i} w^{(i)} \delta(\mathbf{x} - \mathbf{x}^{(i)}).$$

- Approximates Bayesian optimal filtering equations with importance sampling.
- Particle filtering = Sequential importance sampling, with additional resampling step.

Monte Carlo Integration

 In Bayesian inference we often want to compute posterior expectations of the form

$$\mathsf{E}[\mathbf{g}(\mathbf{x})\,|\,\mathbf{y}_{1:T}] = \int \mathbf{g}(\mathbf{x})\; \rho(\mathbf{x}\,|\,\mathbf{y}_{1:T})\;\mathrm{d}\mathbf{x}$$

• Monte Carlo: draw N independent random samples from $\mathbf{x}^{(i)} \sim p(\mathbf{x} \mid \mathbf{y}_{1:T})$ and estimate the expectation as

$$\mathsf{E}[\mathbf{g}(\mathbf{x})\,|\,\mathbf{y}_{1:T}] \approx \frac{1}{N} \sum_{i=1}^{N} \mathbf{g}(\mathbf{x}^{(i)}).$$

Importance Sampling: Basic Version [1/2]

- In practice, we rarely can directly draw samples from the distribution $p(\mathbf{x} | \mathbf{y}_{1:T})$.
- In importance sampling (IS), we draw samples from an importance distribution $\mathbf{x}^{(i)} \sim \pi(\mathbf{x} \mid \mathbf{y}_{1:T})$ and compute weights $\tilde{w}^{(i)}$ such that

$$\mathsf{E}[\mathbf{g}(\mathbf{x})\,|\,\mathbf{y}_{1:T}] \approx \sum_{i=1}^N \tilde{w}^{(i)}\,\mathbf{g}(\mathbf{x}^{(i)})$$

Importance Sampling: Basic Version [2/2]

Importance sampling is based on the identity

$$\begin{aligned} \mathsf{E}[\mathbf{g}(\mathbf{x}) \,|\, \mathbf{y}_{1:T}] &= \int \mathbf{g}(\mathbf{x}) \, \rho(\mathbf{x} \,|\, \mathbf{y}_{1:T}) \, \mathrm{d}\mathbf{x} \\ &= \int \left[\mathbf{g}(\mathbf{x}) \, \frac{\rho(\mathbf{x} \,|\, \mathbf{y}_{1:T})}{\pi(\mathbf{x} \,|\, \mathbf{y}_{1:T})} \right] \, \pi(\mathbf{x} \,|\, \mathbf{y}_{1:T}) \, \mathrm{d}\mathbf{x} \end{aligned}$$

• Thus we can form a Monte Carlo approximation as follows:

$$\mathsf{E}[\mathbf{g}(\mathbf{x}) \,|\, \mathbf{y}_{1:T}] \approx \frac{1}{N} \sum_{i=1}^{N} \frac{\rho(\mathbf{x}^{(i)} \,|\, \mathbf{y}_{1:T})}{\pi(\mathbf{x}^{(i)} \,|\, \mathbf{y}_{1:T})} \,\mathbf{g}(\mathbf{x}^{(i)})$$

That is, the importance weights can be defined as

$$\tilde{w}^{(i)} = \frac{1}{N} \frac{\rho(\mathbf{x}^{(i)} \mid \mathbf{y}_{1:T})}{\pi(\mathbf{x}^{(i)} \mid \mathbf{y}_{1:T})}$$

Importance Sampling: Weight Normalization

- The problem is that we need to evaluate the normalization constant of $p(\mathbf{x}^{(i)} | \mathbf{y}_{1:T})$ often not possible.
- However, it turns out that we get a valid algorithm if we define unnormalized importance weights as

$$w^{*(i)} = \frac{\rho(\mathbf{y}_{1:T} \mid \mathbf{x}^{(i)}) \, \rho(\mathbf{x}^{(i)})}{\pi(\mathbf{x}^{(i)} \mid \mathbf{y}_{1:T})}$$

and then normalize them:

$$\mathbf{w}^{(i)} = \frac{\mathbf{w}^{*(i)}}{\sum_{j} \mathbf{w}^{*(j)}}$$

 The (weight-normalized) importance sampling approximation is then

$$\mathsf{E}[\mathbf{g}(\mathbf{x})\,|\,\mathbf{y}_{1:T}] \approx \sum_{i=1}^N w^{(i)}\,\mathbf{g}(\mathbf{x}^{(i)})$$

Importance Sampling: Algorithm

Importance Sampling

• Draw *N* samples from the importance distribution:

$$\mathbf{x}^{(i)} \sim \pi(\mathbf{x} \mid \mathbf{y}_{1:T}), \qquad i = 1, \dots, N.$$

Compute the unnormalized weights by

$$w^{*(i)} = \frac{p(\mathbf{y}_{1:T} \mid \mathbf{x}^{(i)}) p(\mathbf{x}^{(i)})}{\pi(\mathbf{x}^{(i)} \mid \mathbf{y}_{1:T})},$$

and the normalized weights by

$$w^{(i)} = \frac{w^{*(i)}}{\sum_{j=1}^{N} w^{*(j)}}.$$

Importance Sampling: Properties

• The approximation to the posterior expectation of g(x) is

$$\mathsf{E}[\mathbf{g}(\mathbf{x})\,|\,\mathbf{y}_{1:T}] \approx \sum_{i=1}^N w^{(i)}\,\mathbf{g}(\mathbf{x}^{(i)}).$$

 The posterior probability density approximation can be formally written as

$$p(\mathbf{x} \mid \mathbf{y}_{1:T}) \approx \sum_{i=1}^{N} w^{(i)} \delta(\mathbf{x} - \mathbf{x}^{(i)}),$$

where $\delta(\cdot)$ is the Dirac delta function.

 The efficiency depends on the choice of the importance distribution.

Sequential Importance Sampling: Idea

 Sequential Importance Sampling (SIS) is concerned with models

$$\mathbf{x}_k \sim p(\mathbf{x}_k \mid \mathbf{x}_{k-1})$$

 $\mathbf{y}_k \sim p(\mathbf{y}_k \mid \mathbf{x}_k)$

• The SIS algorithm uses a weighted set of particles $\{(w_k^{(i)}, \mathbf{x}_k^{(i)}) : i = 1, ..., N\}$ such that

$$\mathsf{E}[\mathbf{g}(\mathbf{x}_k) \,|\, \mathbf{y}_{1:k}] \approx \sum_{i=1}^N w_k^{(i)} \mathbf{g}(\mathbf{x}_k^{(i)}).$$

Or equivalently

$$\rho(\mathbf{x}_k \mid \mathbf{y}_{1:k}) \approx \sum_{i=1}^N w_k^{(i)} \delta(\mathbf{x}_k - \mathbf{x}_k^{(i)}),$$

where $\delta(\cdot)$ is the Dirac delta function.

Uses importance sampling sequentially.

Sequential Importance Sampling: Derivation [1/2]

- Let's consider the full posterior distribution of states $\mathbf{x}_{0:k}$ given the measurements $\mathbf{y}_{1:k}$.
- We get the following recursion for the posterior distribution:

$$\begin{aligned} \rho(\mathbf{x}_{0:k} \,|\, \mathbf{y}_{1:k}) &\propto \rho(\mathbf{y}_k \,|\, \mathbf{x}_{0:k}, \mathbf{y}_{1:k-1}) \, \rho(\mathbf{x}_{0:k} \,|\, \mathbf{y}_{1:k-1}) \\ &= \rho(\mathbf{y}_k \,|\, \mathbf{x}_k) \, \rho(\mathbf{x}_k \,|\, \mathbf{x}_{0:k-1}, \mathbf{y}_{1:k-1}) \, \rho(\mathbf{x}_{0:k-1} \,|\, \mathbf{y}_{1:k-1}) \\ &= \rho(\mathbf{y}_k \,|\, \mathbf{x}_k) \, \rho(\mathbf{x}_k \,|\, \mathbf{x}_{k-1}) \, \rho(\mathbf{x}_{0:k-1} \,|\, \mathbf{y}_{1:k-1}). \end{aligned}$$

• We could now construct an importance distribution $\mathbf{x}_{0:k}^{(i)} \sim \pi(\mathbf{x}_{0:k} \mid \mathbf{y}_{1:k})$ and compute the corresponding (normalized) importance weights as

$$w_k^{(i)} \propto \frac{p(\mathbf{y}_k \,|\, \mathbf{x}_k^{(i)}) \, p(\mathbf{x}_k^{(i)} \,|\, \mathbf{x}_{k-1}^{(i)}) \, p(\mathbf{x}_{0:k-1}^{(i)} \,|\, \mathbf{y}_{1:k-1})}{\pi(\mathbf{x}_{0:k}^{(i)} \,|\, \mathbf{y}_{1:k})}.$$

Sequential Importance Sampling: Derivation [2/2]

 Let's form the importance distribution recursively as follows:

$$\pi(\mathbf{x}_{0:k} | \mathbf{y}_{1:k}) = \pi(\mathbf{x}_k | \mathbf{x}_{0:k-1}, \mathbf{y}_{1:k}) \pi(\mathbf{x}_{0:k-1} | \mathbf{y}_{1:k-1})$$

• Expression for the importance weights can be written as

$$w_k^{(i)} \propto \frac{p(\mathbf{y}_k \,|\, \mathbf{x}_k^{(i)}) \, p(\mathbf{x}_k^{(i)} \,|\, \mathbf{x}_{k-1}^{(i)})}{\pi(\mathbf{x}_k^{(i)} \,|\, \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k})} \underbrace{\frac{p(\mathbf{x}_{0:k-1}^{(i)} \,|\, \mathbf{y}_{1:k-1})}{\pi(\mathbf{x}_{0:k-1}^{(i)} \,|\, \mathbf{y}_{1:k-1})}}_{\propto w_{k-1}^{(i)}}$$

Thus the weights satisfy the recursion

$$w_k^{(i)} \propto \frac{p(\mathbf{y}_k \,|\, \mathbf{x}_k^{(i)}) \, p(\mathbf{x}_k^{(i)} \,|\, \mathbf{x}_{k-1}^{(i)})}{\pi(\mathbf{x}_k^{(i)} \,|\, \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k})} \, w_{k-1}^{(i)}$$

Sequential Importance Sampling: Algorithm

Sequential Importance Sampling

• Initialization: Draw N samples $\mathbf{x}_0^{(i)}$ from the prior

$$\mathbf{x}_0^{(i)} \sim p(\mathbf{x}_0)$$

and set $w_0^{(i)} = 1/N$.

Prediction: Draw N new samples x_k⁽ⁱ⁾ from importance distributions

$$\mathbf{x}_{k}^{(i)} \sim \pi(\mathbf{x}_{k} \,|\, \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k})$$

• Update: Calculate new weights according to

$$w_k^{(i)} \propto w_{k-1}^{(i)} rac{
ho(\mathbf{y}_k \,|\, \mathbf{x}_k^{(i)}) \,
ho(\mathbf{x}_k^{(i)} \,|\, \mathbf{x}_{k-1}^{(i)})}{\pi(\mathbf{x}_k^{(i)} \,|\, \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k})}$$

Sequential Importance Sampling: Degeneracy

- The problem in SIS is that the algorithm is degenerate
- It can be shown that the variance of the weights increases at every step
- It means that we will always converge to single non-zero weight $w^{(i)} = 1$ and the rest being zero not very useful algorithm.
- Solution: resampling!

Sequential Importance Resampling: Resampling Step

 Sequential Importance Resampling (SIR) algorithm adds the following resampling step to SIS algorithm:

Resampling

- Interpret each weight $w_k^{(i)}$ as the probability of obtaining the sample index i in the set $\{\mathbf{x}_k^{(i)} \mid i = 1, ..., N\}$.
- Draw *N* samples from that discrete distribution and replace the old sample set with this new one.
- Set all weights to the constant value $w_k^{(i)} = 1/N$.
- There are many algorithms for implementing this stratified resampling is optimal in terms of variance.

Sequential Importance Resampling: Effective Number of Particles

- A simple way to do resampling is at every step but every resampling operation increases variance.
- We can also resample at, say, every Kth step.
- In adaptive resampling, we resample when the effective number of samples is too low (say, N/10):

$$n_{\text{eff}} pprox rac{1}{\sum_{i=1}^{N} \left(w_k^{(i)}\right)^2},$$

 In theory, biased, but in practice works very well and is often used.

Sequential Importance Resampling: Algorithm

Sequential Importance Resampling

• Draw point $\mathbf{x}_k^{(i)}$ from the importance distribution:

$$\mathbf{x}_{k}^{(i)} \sim \pi(\mathbf{x}_{k} \mid \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k}), \qquad i = 1, \dots, N.$$

Calculate new weights

$$w_k^{(i)} \propto w_{k-1}^{(i)} \frac{\rho(\mathbf{y}_k \mid \mathbf{x}_k^{(i)}) \ \rho(\mathbf{x}_k^{(i)} \mid \mathbf{x}_{k-1}^{(i)})}{\pi(\mathbf{x}_k^{(i)} \mid \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k})}, \qquad i = 1, \dots, N,$$

and normalize them to sum to unity.

 If the effective number of particles is too low, perform resampling.

Sequential Importance Resampling: Bootstrap filter

 In bootstrap filter we use the dynamic model as the importance distribution

$$\pi(\mathbf{x}_{k}^{(i)} \mid \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k}) = p(\mathbf{x}_{k}^{(i)} \mid \mathbf{x}_{k-1}^{(i)})$$

and resample at every step:

Bootstrap Filter

• Draw point $\mathbf{x}_k^{(i)}$ from the dynamic model:

$$\mathbf{x}_k^{(i)} \sim p(\mathbf{x}_k \mid \mathbf{x}_{k-1}^{(i)}), \qquad i = 1, \dots, N.$$

Calculate new weights

$$w_k^{(i)} \propto p(\mathbf{y}_k \mid \mathbf{x}_k^{(i)}), \qquad i = 1, \ldots, N,$$

and normalize them to sum to unity.

Perform resampling.

Sequential Importance Resampling: Optimal Importace Distribution

• The optimal importance distribution is

$$\pi(\mathbf{x}_{k}^{(i)} \mid \mathbf{x}_{0:k-1}^{(i)}, \mathbf{y}_{1:k}) = p(\mathbf{x}_{k}^{(i)} \mid \mathbf{x}_{k-1}^{(i)}, \mathbf{y}_{k})$$

• Then the weight update reduces to

$$w_k^{(i)} \propto w_{k-1}^{(i)} \, \rho(\mathbf{y}_k \mid \mathbf{x}_{k-1}^{(i)}), \qquad i = 1, \dots, N.$$

 The optimal importance distribution can be used, for example, when the state space is finite.

Sequential Importance Resampling: Importace Distribution via Kalman Filtering

 We can also form a Gaussian approximation to the optimal importance distribution:

$$p(\mathbf{x}_k^{(i)} \mid \mathbf{x}_{k-1}^{(i)}, \mathbf{y}_k) \approx \mathsf{N}(\mathbf{x}_k^{(i)} \mid \tilde{\mathbf{m}}_k^{(i)}, \tilde{\mathbf{P}}_k^{(i)}).$$

by using a single prediction and update steps of a Gaussian filter starting from a singular distribution at $\mathbf{x}_{k-1}^{(i)}$.

- We can also replace above with the result of a Gaussian filter $N(\mathbf{m}_{k-1}^{(i)}, \mathbf{P}_{k-1}^{(i)})$ started from a random initial mean.
- A very common way seems to be to use the previous sample as the mean: $N(\mathbf{x}_{k-1}^{(i)}, \mathbf{P}_{k-1}^{(i)})$.
- A particle filter with UKF proposal has been given name unscented particle filter (UPF) – you can invent new PFs easily this way.

Rao-Blackwellized Particle Filter: Idea

 Rao-Blackwellized particle filtering (RBPF) is concerned with conditionally Gaussian models:

$$\begin{split} \rho(\mathbf{x}_{k} \,|\, \mathbf{x}_{k-1}, \mathbf{u}_{k-1}) &= \mathsf{N}(\mathbf{x}_{k} \,|\, \mathbf{A}_{k-1}(\mathbf{u}_{k-1}) \,\mathbf{x}_{k-1}, \mathbf{Q}_{k-1}(\mathbf{u}_{k-1})) \\ \rho(\mathbf{y}_{k} \,|\, \mathbf{x}_{k}, \mathbf{u}_{k}) &= \mathsf{N}(\mathbf{y}_{k} \,|\, \mathbf{H}_{k}(\mathbf{u}_{k}) \,\mathbf{x}_{k}, \mathbf{R}_{k}(\mathbf{u}_{k})) \\ \rho(\mathbf{u}_{k} \,|\, \mathbf{u}_{k-1}) &= (\mathsf{any} \; \mathsf{given} \; \mathsf{form}), \end{split}$$

where

- \mathbf{x}_k is the state
- y_k is the measurement
- **u**_k is an arbitrary latent variable
- Given the latent variables $\mathbf{u}_{1:T}$ the model is Gaussian.
- The RBPF uses SIR for the latent variables and computes the conditionally Gaussian part in closed form with Kalman filter.

Rao-Blackwellized Particle Filter: Derivation [1/3]

• The full posterior at step *k* can be factored as

$$p(\mathbf{u}_{0:k}, \mathbf{x}_{0:k} \mid \mathbf{y}_{1:k}) = p(\mathbf{x}_{0:k} \mid \mathbf{u}_{0:k}, \mathbf{y}_{1:k}) p(\mathbf{u}_{0:k} \mid \mathbf{y}_{1:k})$$

- The first term is Gaussian and computable with Kalman filter and RTS smoother
- For the second term we get the following recursion:

$$\begin{split} & \rho(\mathbf{u}_{0:k} \mid \mathbf{y}_{1:k}) \\ & \propto \rho(\mathbf{y}_k \mid \mathbf{u}_{0:k}, \mathbf{y}_{1:k-1}) \, \rho(\mathbf{u}_{0:k} \mid \mathbf{y}_{1:k-1}) \\ & = \rho(\mathbf{y}_k \mid \mathbf{u}_{0:k}, \mathbf{y}_{1:k-1}) \, \rho(\mathbf{u}_k \mid \mathbf{u}_{0:k-1}, \mathbf{y}_{1:k-1}) \, \rho(\mathbf{u}_{0:k-1} \mid \mathbf{y}_{1:k-1}) \\ & = \rho(\mathbf{y}_k \mid \mathbf{u}_{0:k}, \mathbf{y}_{1:k-1}) \, \rho(\mathbf{u}_k \mid \mathbf{u}_{k-1}) \, \rho(\mathbf{u}_{0:k-1} \mid \mathbf{y}_{1:k-1}) \end{split}$$

Rao-Blackwellized Particle Filter: Derivation [2/3]

Let's take a look at the terms in

$$p(\mathbf{y}_k | \mathbf{u}_{0:k}, \mathbf{y}_{1:k-1}) p(\mathbf{u}_k | \mathbf{u}_{k-1}) p(\mathbf{u}_{0:k-1} | \mathbf{y}_{1:k-1})$$

- The first term can be computed by running Kalman filter with fixed u_{0:k} over the measurement sequence.
- The second term is just the dynamic model.
- The third term is the posterior from the previous step.

Rao-Blackwellized Particle Filter: Derivation [3/3]

• We can form the importance distribution recursively:

$$\pi(\mathbf{u}_{0:k} \mid \mathbf{y}_{1:k}) = \pi(\mathbf{u}_k \mid \mathbf{u}_{0:k-1}, \mathbf{y}_{1:k}) \pi(\mathbf{u}_{0:k-1} \mid \mathbf{y}_{1:k-1})$$

• We then get the following recursion for the weights:

$$w_k^{(i)} \propto \frac{\rho(\mathbf{y}_k \mid \mathbf{u}_{0:k-1}^{(i)}, \mathbf{y}_{1:k-1}) \, \rho(\mathbf{u}_k^{(i)} \mid \mathbf{u}_{k-1}^{(i)})}{\pi(\mathbf{u}_k^{(i)} \mid \mathbf{u}_{0:k-1}^{(i)}, \mathbf{y}_{1:k})} \, w_{k-1}^{(i)}$$

- Given the marginal posterior for u_{0:k} we can recover the Gaussian part x_{0:k} with Kalman filter and RTS smoother.
- The optimal importance distribution takes the form

$$p(\mathbf{u}_k \mid \mathbf{y}_{1:k}, \mathbf{u}_{0:k-1}^{(i)}) \propto p(\mathbf{y}_k \mid \mathbf{u}_k, \mathbf{u}_{0:k-1}^{(i)}) p(\mathbf{u}_k \mid \mathbf{u}_{0:k-1}^{(i)}, \mathbf{y}_{1:k-1})$$

Rao-Blackwellized Particle Filter: Algorithm [1/3]

Rao-Blackwellized Particle Filter

• Perform Kalman filter predictions for each of the Kalman filter means and covariances in the particles i = 1, ..., N conditional on the previously drawn latent variable values $\mathbf{u}_{k-1}^{(i)}$

$$\begin{split} & \boldsymbol{m}_k^{-(i)} = \boldsymbol{A}_{k-1}(\boldsymbol{u}_{k-1}^{(i)}) \, \boldsymbol{m}_{k-1}^{(i)} \\ & \boldsymbol{P}_k^{-(i)} = \boldsymbol{A}_{k-1}(\boldsymbol{u}_{k-1}^{(i)}) \, \boldsymbol{P}_{k-1}^{(i)} \, \boldsymbol{A}_{k-1}^T(\boldsymbol{u}_{k-1}^{(i)}) + \boldsymbol{Q}_{k-1}(\boldsymbol{u}_{k-1}^{(i)}). \end{split}$$

• Draw new latent variables $\mathbf{u}_k^{(i)}$ for each particle in i = 1, ..., N from the corresponding importance distributions

$$\mathbf{u}_{k}^{(i)} \sim \pi(\mathbf{u}_{k} \mid \mathbf{u}_{0:k-1}^{(i)}, \mathbf{y}_{1:k}).$$

Rao-Blackwellized Particle Filter: Algorithm [2/3]

Rao-Blackwellized Particle Filter (cont.)

Calculate new weights as follows:

$$w_k^{(i)} \propto w_{k-1}^{(i)} \frac{\rho(\mathbf{y}_k \mid \mathbf{u}_{0:k}^{(i)}, \mathbf{y}_{1:k-1}) \ \rho(\mathbf{u}_k^{(i)} \mid \mathbf{u}_{k-1}^{(i)})}{\pi(\mathbf{u}_k^{(i)} \mid \mathbf{u}_{0:k-1}^{(i)}, \mathbf{y}_{1:k})},$$

where the likelihood term is the marginal measurement likelihood of the Kalman filter:

$$\rho(\mathbf{y}_{k} \mid \mathbf{u}_{0:k}^{(i)}, \mathbf{y}_{1:k-1}) \\
= N\left(\mathbf{y}_{k} \mid \mathbf{H}_{k}(\mathbf{u}_{k}^{(i)}) \, \mathbf{m}_{k}^{-(i)}, \mathbf{H}_{k}(\mathbf{u}_{k}^{(i)}) \, \mathbf{P}_{k}^{-(i)} \, \mathbf{H}_{k}^{T}(\mathbf{u}_{k}^{(i)}) + \mathbf{R}_{k}(\mathbf{u}_{k}^{(i)})\right).$$

Then normalize the weights to sum to unity.

Rao-Blackwellized Particle Filter: Algorithm [3/3]

Rao-Blackwellized Particle Filter (cont.)

• Perform Kalman filter updates for each of the particles conditional on the drawn latent variables $\mathbf{u}_k^{(i)}$

$$\mathbf{v}_{k}^{(i)} = \mathbf{y}_{k} - \mathbf{H}_{k}(\mathbf{u}_{k}^{(i)}) \, \mathbf{m}_{k}^{-}$$

$$\mathbf{S}_{k}^{(i)} = \mathbf{H}_{k}(\mathbf{u}_{k}^{(i)}) \, \mathbf{P}_{k}^{-(i)} \, \mathbf{H}_{k}^{T}(\mathbf{u}_{k}^{(i)}) + \mathbf{R}_{k}(\mathbf{u}_{k}^{(i)})$$

$$\mathbf{K}_{k}^{(i)} = \mathbf{P}_{k}^{-(i)} \, \mathbf{H}_{k}^{T}(\mathbf{u}_{k}^{(i)}) \, \mathbf{S}_{k}^{-1}$$

$$\mathbf{m}_{k}^{(i)} = \mathbf{m}_{k}^{-(i)} + \mathbf{K}_{k}^{(i)} \, \mathbf{v}_{k}^{(i)}$$

$$\mathbf{P}_{k}^{(i)} = \mathbf{P}_{k}^{-(i)} - \mathbf{K}_{k}^{(i)} \, \mathbf{S}_{k}^{(i)} \, [\mathbf{K}_{k}^{(i)}]^{T}.$$

 If the effective number of particles is too low, perform resampling.

Rao-Blackwellized Particle Filter: Properties

- The Rao-Blackwellized particle filter produces a set of weighted samples $\{w_k^{(i)}, \mathbf{u}_k^{(i)}, \mathbf{m}_k^{(i)}, \mathbf{P}_k^{(i)} : i = 1, ..., N\}$
- The expectation of a function $g(\cdot)$ can be approximated as

$$\mathsf{E}[\mathbf{g}(\mathbf{x}_k, \mathbf{u}_k) \,|\, \mathbf{y}_{1:k}] \approx \sum_{i=1}^N w_k^{(i)} \, \int \mathbf{g}(\mathbf{x}_k, \mathbf{u}_k^{(i)}) \, \, \mathsf{N}(\mathbf{x}_k \,|\, \mathbf{m}_k^{(i)}, \mathbf{P}_k^{(i)}) \, \mathrm{d}\mathbf{x}_k.$$

Approximation of the filtering distribution is

$$p(\mathbf{x}_k, \mathbf{u}_k | \mathbf{y}_{1:k}) \approx \sum_{i=1}^N w_k^{(i)} \, \delta(\mathbf{u}_k - \mathbf{u}_k^{(i)}) \, N(\mathbf{x}_k | \mathbf{m}_k^{(i)}, \mathbf{P}_k^{(i)}).$$

• It is possible to do approximate Rao-Blackwellization by replacing the Kalman filter with a Gaussian filter.

Rao-Blackwellization of Static Parameters

 Rao-Blackwellization can sometimes be used in models of the form

$$egin{aligned} \mathbf{x}_k &\sim p(\mathbf{x}_k \,|\, \mathbf{x}_{k-1}, \mathbf{u}) \ \mathbf{y}_k &\sim p(\mathbf{y}_k \,|\, \mathbf{x}_k, \mathbf{u}) \ \mathbf{u} &\sim p(\mathbf{u}), \end{aligned}$$

where vector **u** contains the unknown static parameters.

 Possible if the posterior distribution of parameters u depends only on some sufficient statistics T_k:

$$\mathbf{T}_k = \mathbf{T}_k(\mathbf{x}_{1:k}, \mathbf{y}_{1:k})$$

- We also need to have a recursion rule for the sufficient statistics.
- Can be extended to time-varying parameters.

Particle Filter: Advantages

- No restrictions in model can be applied to non-Gaussian models, hierarchical models etc.
- Global approximation.
- Approaches the exact solution, when the number of samples goes to infinity.
- In its basic form, very easy to implement.
- Superset of other filtering methods Kalman filter is a Rao-Blackwellized particle filter with one particle.

Particle Filter: Disadvantages

- Computational requirements much higher than of the Kalman filters.
- Problems with nearly noise-free models, especially with accurate dynamic models.
- Good importance distributions and efficient Rao-Blackwellized filters quite tricky to implement.
- Very hard to find programming errors (i.e., to debug).

Summary

- Particle filters use weighted set of samples (particles) for approximating the filtering distributions.
- Sequential importance resampling (SIR) is the general framework and bootstrap filter is a simple special case of it.
- EKF, UKF and other Gaussian filters can be used for forming good importance distributions.
- In Rao-Blackwellized particle filters a part of the state is sampled and part is integrated in closed form with Kalman filter.

Particle Filter: Demo

• The discretized pendulum model:

$$\begin{pmatrix} x_k^1 \\ x_k^2 \end{pmatrix} = \underbrace{\begin{pmatrix} x_{k-1}^1 + x_{k-1}^2 \Delta t \\ x_{k-1}^2 - g \sin(x_{k-1}^1) \Delta t \end{pmatrix}}_{\mathbf{f}(\mathbf{x}_{k-1})} + \begin{pmatrix} 0 \\ q_{k-1} \end{pmatrix}$$

$$y_k = \underbrace{\sin(x_k^1)}_{\mathbf{h}(\mathbf{x}_k)} + r_k,$$

→ Matlab demonstration